

Lecture 16

Discrete Choice Models
Limited Dependent Models

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Lecture 16: Objectives

- ◆ Discrete Choice Models
 - Binary Choice (lecture 15)
 - Multiple Choice
 - ◆ Unordered choice sets
 - ◆ Ordered choice sets
- ◆ Count Data
- ◆ Limited Dependent Models
 - Tobit Models
 - Sample Selection

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Lecture 16: Multiple Choice

- ◆ Multiple Choice:
One decision; Many choices
 - Unordered choice sets
 - ◆ Urban transport (car, train, bus, bike)
 - ◆ Occupational choice
 - Ordered choice sets
 - ◆ Bond ratings
 - ◆ Exam grades
- ◆ Multivariate Choice: Several decisions
 - Where to live (city/suburb); type of housing (house/flat)

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Unordered Choice Sets

- ◆ Choices: $j = 1, 2, \dots, J$
 We want to model $P(Y = j)$
 Explanatory variables: $z_{ij} = [x_{ij} \ w_i]$
 x_{ij} : choice attributes
 w_i : individual characteristics
- ◆ Define an index function: $\theta_{ij} = x'_{ij}\beta + w'_i\gamma_j$
 Maximising a random utility $U_{ij} = \theta_{ij} + \varepsilon_{ij}$
 across choices allows us to model $P(Y_i = j)$
- ◆ Probit models get complicated if $J > 2$

Multinomial Logit

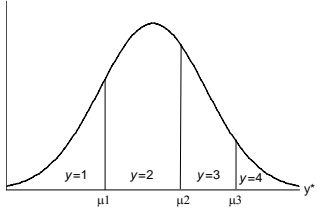
- ◆ Most work is done using logit models
- $$P(Y_i = j) = \frac{e^{\theta_{ij}}}{\sum_{j=1}^J e^{\theta_{ij}}} = \frac{e^{x'_{ij}\beta + w'_i\gamma_j}}{\sum_{j=1}^J e^{x'_{ij}\beta + w'_i\gamma_j}}$$
- ◆ Increasing every γ with a constant doesn't affect this probability. We normalise $\gamma_1 = 0$
 - ◆ We are comparing choices with a baseline decision, in this case choice #1

Multinomial Logit, cont

- ◆ Some authors distinguish between *multinomial models* ($\beta = 0$), *conditional models* ($\gamma = 0$) and *mixed models* ($\beta, \gamma \neq 0$)
- ◆ These models can all be estimated using maximum likelihood
- ◆ Logit models imply that ratios of probabilities are independent of other choices:
IIA (independence of irrelevant alternatives).
 Probit or nested logit models can be used

Ordered Choice Sets

- The ordered probit (or logit) uses a latent variable representation

$$y_i^* = \mathbf{x}_i' \boldsymbol{\beta} + \varepsilon_i$$


$$P(Y_i = j) = P(\mu_{j-1} < y_i^* < \mu_j)$$

- $\boldsymbol{\beta}$ and $\boldsymbol{\mu}$ can be estimated using maximum likelihood

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Count Data

- In some cases the numbers we assign to the “choices” mean something. That is, y can be interpreted as a discrete random variable.
- We can use OLS regression (normal QML)
- COUNT DATA** models are to be preferred if y comes from so called “rare events” with many zeroes and small values. For example:
 - number of patents registered in a period
 - number of sick days per year

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The Poisson Model

- The simplest count model assumes that no more than one event can occur in a short period of time, and that there is independence between these time periods

⇒ the **POISSON** model: $E(y_i | \mathbf{x}_i) = e^{\mathbf{x}_i' \boldsymbol{\beta}}$

where $y | \mathbf{x}$ has a Poisson distribution

- This is usually estimated by ML from

$$\ln L = \sum [-e^{\mathbf{x}_i' \boldsymbol{\beta}} + y_i \mathbf{x}_i' \boldsymbol{\beta} - \ln y_i!]$$

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Problems with Poisson models

- ◆ The Poisson distribution has variance equal to mean. In empirical studies we often observe **overdispersion**: (variance > mean)
- ◆ There are several tests for overdispersion
- ◆ Robust standard errors exist (Poisson QML)
- ◆ We can extend the model to account for heterogeneity and/or dependence. One example is the **NEGATIVE BINOMIAL** model

Censoring vs Truncation

- ◆ Difference between censoring and truncation
 - Example: 5 shots of which 4 hit the target
 - ◆ Censoring: misses are recorded as 0
 - ◆ Truncating: only shots on target are recorded
- ◆ Modeling y with the aid of \mathbf{x}
 - CENSORING:** y_i is measured only if $y_i > \tau$
 \mathbf{x}_i is measured for all i
 - TRUNCATION:** \mathbf{x}_i and y_i measured only if $y_i > \tau$

Truncated Regression

- ◆ A truncated regression $y_i = \mathbf{x}'_i \boldsymbol{\beta} + \varepsilon_i$ exists for all i but is only observed for $y_i > \tau$
- ◆ $E(y_i | \mathbf{x}_i, y_i > \tau) > E(y_i | \mathbf{x}_i) = \mathbf{x}'_i \boldsymbol{\beta}$ since we are excluding small values of y
- ◆ It can be shown that

$$E(y_i | \mathbf{x}_i, y_i > \tau) = \mathbf{x}'_i \boldsymbol{\beta} + \sigma \lambda((\tau - \mathbf{x}'_i \boldsymbol{\beta}) / \sigma)$$

$$\lambda(z) = \frac{f(z)}{1 - F(z)} \quad \text{INVERSE MILLS RATIO}$$
 where f and F are the p.d.f. and c.d.f. of ε

Truncated Regression, cont

- ◆ The truncated regression can thus be written

$$y_i = \mathbf{x}'_i\beta + \sigma\lambda((\tau - \mathbf{x}'_i\beta)/\sigma) + v_i$$
 where $E(v|\mathbf{x})=0$
- ◆ This can be estimated using ML
- ◆ Truncated regression is rarely used directly, since we usually have observations on \mathbf{x}_i when $y_i \leq \tau$, or τ is stochastic.

Tobit Model: Censored Regression

- ◆ The model: $y_i = \begin{cases} \tau & y_i^* \leq \tau \\ y_i^* & y_i^* > \tau \end{cases}$
 $y_i^* = \mathbf{x}'_i\beta + \varepsilon_i$ with $\varepsilon_i \sim N(0, \sigma^2)$
- ◆ OLS will be inconsistent since

$$E(y_i | \mathbf{x}_i) = P(y_i \leq \tau)E(y_i | \mathbf{x}_i, y_i \leq \tau) + P(y_i > \tau)E(y_i | \mathbf{x}_i, y_i > \tau)$$

$$= \Phi(\alpha)\tau + (1 - \Phi(\alpha))(\mathbf{x}'_i\beta + \sigma\lambda(\alpha))$$
 where $\alpha = (\tau - \mathbf{x}'_i\beta)/\sigma$

Tobit Model; cont

- ◆ The Tobit model is estimated using ML
- ◆ The marginal effects are given by

$$\psi_j = \beta_j\Phi(\mathbf{x}'\beta/\sigma)$$
- ◆ Problems:
 - Heteroscedasticity
 - Non-normality
 - The same model assumed for y^* both above and below the threshold

Sample Selection Models

- ◆ We have two processes.
 Selection Process: $z_i^* = \mathbf{w}'_i \boldsymbol{\gamma} + v_i$
 Regression Process: $y_i = \mathbf{x}'_i \boldsymbol{\beta} + \varepsilon_i$ if $z_i^* > 0$
 (ε, v) is bivariate normally distributed with $\rho = \text{Cor}(\varepsilon, v)$
- ◆ This is sometimes called Tobit type II model
- ◆ Examples
 - Advantages of moving only observed for movers
 - Purchases of seldom bought goods only observed for purchasers

Sample Selection, cont

- ◆ We can estimate with ML, but usually done with Heckman's Two-Step method (Heckit).
- 1) Estimate the probit model

$$\Rightarrow \hat{\lambda}_i = \frac{\phi(-\mathbf{w}'_i \hat{\boldsymbol{\gamma}})}{1 - \Phi(-\mathbf{w}'_i \hat{\boldsymbol{\gamma}})} = \frac{\phi(\mathbf{w}'_i \hat{\boldsymbol{\gamma}})}{\Phi(\mathbf{w}'_i \hat{\boldsymbol{\gamma}})}$$
- 2) Estimate the regression $y_i = \mathbf{x}'_i \boldsymbol{\beta} + (\rho \sigma_\varepsilon) \hat{\lambda}_i + u_i$
- ◆ To identify the model empirically we need some elements of \mathbf{w} not to be present in \mathbf{x}

Sample Selection, cont

- ◆ The standard errors of the Heckit procedure can be calculated in a number of ways. We have to take account of the fact that we are estimating λ
- ◆ The estimated marginal effects are given by

$$\hat{\psi}_j = \hat{\beta}_j [1 - \hat{\lambda}_i (\hat{\lambda}_i - \mathbf{w}'_i \hat{\boldsymbol{\gamma}})]$$
